

Utrecht University
Mathematical Institute

**Answers Re-Examination for Introduction to Financial
Mathematics, WISB373**

Monday March 13th 2023, 17:00-20:00 o'clock (**3 hours examination**)

(Each item is worth 10 points)

1. Assume we have a European call $c(t)$ and a put option $p(t)$, with the same expiry date $T = 4$, i.e., exercise in 4 years, and strike price $K = 10$ Euro. The current share price is 11 Euro, assuming a zero constant interest rate $r = 0\%$. Determine if there exists an arbitrage opportunity if both options currently have the value $c(0) = 2.5$ Euro and $p(0) = 1.5$ Euro.

Answer 1. We form two portfolios using the options, the underlying asset and a cash amount K , with one based on the put $p(t)$ and the other based on the call $c(t)$, as follows,

$$\begin{aligned}\Pi_1(t) &= p(t) + S(t), \\ \Pi_2(t) &= c(t) + K.\end{aligned}$$

(as $e^0 = 1$). These portfolios have same value at expiry time T . moreover, by the put-call parity, we see that their values are also equal any time prior to the exercise time, particularly at time $t = 0$. So, there is no arbitrage opportunity here!

2. a. The random process $Z(t)$ is defined as $Z(t) = \alpha W(t) - \sqrt{\beta} W^*(t)$, where $W(t)$ and $W^*(t)$ are independent standard Brownian motions. Determine the relationship between α and β for which $Z(t)$ is a Brownian motion.
- b. Determine whether $W(t) + 4t$ is a martingale.
- c. Let $v_1, v_2, v_3 \in \mathbb{R}^3$ be orthonormal vectors, i.e. $v_i \cdot v_j = \delta_{ij}$. If $W(t) = (W_1(t), W_2(t), W_3(t))$ is a three-dimensional Brownian motion and $X_j(t) = v_j \cdot W(t)$ for $j \in \{1, 2, 3\}$, show, with the help of Lévy's characterization, that (X_1, X_2, X_3) is another three-dimensional Brownian motion.

Answer 2a.: For the combination of Brownian Motions, we find:

$$\begin{aligned}\mathbb{E}[Z(t)] &= \mathbb{E}[\alpha W(t) - \sqrt{\beta} W^*(t)] \\ &= \alpha \mathbb{E}[W(t)] - \sqrt{\beta} \mathbb{E}[W^*(t)] \\ &= 0\end{aligned}$$

For the variance, we find:

$$\begin{aligned}\text{Var}[Z(t+u) - Z(t)] &= \text{Var}[(\alpha W(t+u) - \sqrt{\beta} W^*(t+u)) \\ &\quad - (\alpha W(t) - \sqrt{\beta} W^*(t))] \\ &= \text{Var}[(\alpha(W(t+u) - W(t)) - \sqrt{\beta}(W^*(t+u) - W^*(t)))] \\ &= \text{Var}[(\alpha(W(t+u) - W(t)))] - \text{Var}[\sqrt{\beta}(W^*(t+u) - W^*(t))]\end{aligned}$$

because $W(t)$ and $W^*(t)$ are independent. So,

$$\text{Var}[Z(t+u) - Z(t)] = \alpha^2 u + \beta u = (\alpha^2 + \beta)u.$$

This expression should equal u and not depend on t . It follows that: $\text{Var}[Z(t+u) - Z(t)] = u$ if $\alpha^2 + \beta = 1$ or $\beta = 1 - \alpha^2$.

Answer 2b.: Let $s < t$. Substituting $W(t) = W(s) + (W(t) - W(s))$, gives

$$\begin{aligned}\mathbb{E}[W(t) + 4t|\mathcal{F}(s)] &= \mathbb{E}[W(s) + (W(t) - W(s)) + 4t|\mathcal{F}(s)] \\ &= \mathbb{E}[W(s)|\mathcal{F}(s)] + \mathbb{E}[W(t) - W(s)|\mathcal{F}(s)] + \mathbb{E}[4t|\mathcal{F}(s)] \\ &= W(s) + 0 + 4t\end{aligned}$$

So, $\mathbb{E}[W(t)+4t|\mathcal{F}(s)] \neq W(s)+4s$. Thus $W(t)+4t$ is not a martingale. There was no need to decompose t .

Answer 2c.: We make use of Lévy's characterisation of Brownian motion: each X_j is a linear combination of continuous martingales (the entries W_j), hence a continuous martingale. Also $X_j(0) = 0$ for all j , since the underlying Brownian motions start at zero. Moreover,

$$\begin{aligned}[X_j, X_j](t) &= \left[\sum_{i=1}^3 v_{ji}W_i(t), \sum_{i=1}^3 v_{ji}W_i(t) \right] = \sum_{i,r=1}^3 v_{ji}v_{jr}[W_i, W_r](t) \\ &= \sum_{i=1}^3 v_{ji}^2 t = \|v_j\|^2 t = t\end{aligned}$$

using the orthonormality property of the v_j . Further, for $j \neq \ell$,

$$\begin{aligned}[X_j, X_\ell](t) &= \sum_{i,r=1}^3 v_{ji}v_{\ell,r}[W_i, W_r](t) \\ &= \sum_{i=1}^3 v_{ji}v_{\ell,i} t = v_j \cdot v_\ell = 0\end{aligned}$$

using the orthonormality property of the v_j .

3 Let $Q(t)$ denote the exchange rate at time t . It is the price in domestic currency of one unit of foreign currency and converts foreign currency into domestic currency. A model for the dynamics of the exchange rate is

$$dQ(t)/Q(t) = \mu_Q dt + \sigma_Q dW(t).$$

This has the same structure as the common model for the stock price. The reverse exchange rate, denoted $R(t)$, is the price in foreign currency of one unit of domestic currency $R(t) = 1/Q(t)$. Derive $dR(t)$

Answer 3: $R = 1/Q$ is a function of a single variable Q , so the Itô's-Doeblin formula says:

$$dR = \frac{dR}{dQ}dQ + \frac{1}{2} \frac{d^2 R}{dQ^2}(dQ)^2.$$

Substituting

$$\begin{aligned}dQ &= Q[\mu_Q dt + \sigma_Q dW], \\ (dQ)^2 &= Q^2 \sigma_Q^2 dt \\ \frac{dR}{dQ} &= \frac{-1}{Q^2} \frac{d^2 R}{dQ^2} = \frac{2}{Q^3}.\end{aligned}$$

gives

$$\begin{aligned}dR &= \frac{-1}{Q^2} Q[\mu_Q dt + \sigma_Q dW] + \frac{1}{2} \frac{2}{Q^3} Q^2 \sigma_Q^2 dt \\ &= \frac{-1}{Q} [\mu_Q dt + \sigma_Q dW] + \frac{1}{Q} \sigma_Q^2 dt \\ &= -R[\mu_Q dt + \sigma_Q dW] + R\sigma_Q^2 dt \\ &= R[-\mu_Q + \sigma_Q^2]dt - R\sigma_Q dW.\end{aligned}$$

Dividing by $R(t) \neq 0$ gives the dynamics of $R(t)$:

$$\frac{dR(t)}{R(t)} = (-\mu_Q + \sigma_Q^2)dt - \sigma_Q dW(t)$$

4. Let $\{W(t) : t \geq 0\}$ be a Brownian motion with filtration $\{\mathcal{F}(t) : t \geq 0\}$.

Let $Y(t) = \int_0^t W^2(u)dW(u) - \frac{1}{2} \int_0^t W^4(u)du$ and $X(t) = e^{Y(t)}$, for $t \geq 0$.

a. Prove that $X(t) = 1 + \int_0^t X(u)W^2(u)dW(u)$, for $t \geq 0$.

b. Prove that the process $\{X(t) : t \geq 0\}$ is a martingale with respect to the filtration $\{\mathcal{F}(t) : t \geq 0\}$. Show that $\mathbb{E}(X(t)) = 1$ and

$$\text{Var}[X(t)] = \int_0^t \mathbb{E}[W^4(u)X^2(u)]du \text{ for } t \geq 0.$$

Answer 4a.: Note that $\{Y(t) : t \geq 0\}$ is an Itô process with $dY(t) = W^2(t)dW(t) - \frac{1}{2}W^4(t)dt$ and $dY(t)dY(t) = W^4(t)dt$. Using the Itô-Doeblin formula for Itô processes with $f(x) = e^x$, we get

$$\begin{aligned} X(t) &= f(Y(t)) \\ &= f(Y(0)) + \int_0^t X(u)dY(u) + \frac{1}{2} \int_0^t X(u)dY(u)dY(u) \\ &= 1 + \int_0^t X(u)W^2(u)dW(u) - \frac{1}{2} \int_0^t X(u)W^4(u)du + \frac{1}{2} \int_0^t X(u)W^4(u)du \\ &= 1 + \int_0^t X(u)W^2(u)dW(u). \end{aligned}$$

Answer 4b.: First note that the process $\{Y(t) : t \geq 0\}$ is an Itô process. Since the Itô integral $\{\int_0^t X(u)W^2(u)dW(u) : t \geq 0\}$ seen as a process is a martingale, we conclude that the process $\{X(t) : t \geq 0\}$ is a martingale with respect to the filtration $\{\mathcal{F}(t) : t \geq 0\}$. Thus, $\mathbb{E}[X(t)] = \mathbb{E}[X(0)] = 1$. To calculate the variance, we use the Itô-isometry and Fubini's Theorem (to interchange the integral with the expectation) to get

$$\begin{aligned} \text{Var}[X(t)] &= \mathbb{E}[(X(t) - 1)^2] \\ &= \mathbb{E} \left[\left(\int_0^t X(u)W^2(u)dW(u) \right)^2 \right] \\ &= \mathbb{E} \left[\int_0^t X^2(u)W^4(u)du \right] \\ &= \int_0^t \mathbb{E}[X^2(u)W^4(u)]du. \end{aligned}$$

5. Given a Radon-Nikodym derivative Z , and the associated Radon-Nikodym process $\{Z(t) : t \geq 0\}$, defined by $Z(t) = \mathbb{E}[Z|\mathcal{F}(t)]$, where $\{\mathcal{F}(t) : t \geq 0\}$ is a given filtration. We then have the change of probability measure, $d\tilde{\mathbb{P}} = Zd\mathbb{P}$, with the expectation under the $\tilde{\mathbb{P}}$ -measure, i.e., $\tilde{\mathbb{E}}[Y] = \mathbb{E}[ZY]$

Let Y be a random variable which is $\mathcal{F}(t)$ -measurable. Prove that

$$\mathbb{E}[YZ] = \tilde{\mathbb{E}}[Y] = \mathbb{E}[YZ(t)].$$

Suppose Y is $\mathcal{F}(t)$ -measurable, then prove (using partial averaging) that, for $s < t$

$$\tilde{\mathbb{E}}[Y|\mathcal{F}(s)] = \frac{1}{Z(s)} \mathbb{E}[YZ(t)|\mathcal{F}(s)].$$

Answer 5.: We look for the proof of Lemma 5.2.1 from the book. Recall: $\tilde{\mathbb{E}}[Y] = \mathbb{E}[ZY]$, Y is a r.v. so Y is $\mathcal{F}(t)$ -measurable.

Let $\{\mathcal{F}(t) : t \geq 0\}$ be a given filtration (for which we have defined the Radon-Nikodym process).

We consider the RHS

$$\mathbb{E}[YZ(t)] = \mathbb{E}[Y\mathbb{E}[Z|\mathcal{F}(t)]] = \mathbb{E}[\mathbb{E}[YZ|\mathcal{F}(t)]] = \mathbb{E}[YZ] = \tilde{\mathbb{E}}[Y]$$

using the $\mathcal{F}(t)$ measurability of Y .

Here we look for the proof of book's Lemma 5.2.2:

Recall: $\tilde{\mathbb{E}}[Y] = \mathbb{E}[ZY]$, Y is a r.v., so Y is $\mathcal{F}(t)$ -measurable. To prove the result, it is enough to show that the RHS is the conditional expectation of Y given $\mathcal{F}(s)$ under the measure $\tilde{\mathbb{P}}$.

So, we need to verify the two defining conditions of conditional expectations.

(i) Clearly the RHS is $\mathcal{F}(s)$ -measurable. $Z(s)^{-1}$ is $\mathcal{F}(s)$ -measurable; the same holds for the second term. Hence, the product is $\mathcal{F}(s)$ -measurable.

(ii) Now, let $A \in \mathcal{F}(s)$, we want to show

$$\begin{aligned} \int_A \frac{1}{Z(s)} \mathbb{E}[YZ(t)|\mathcal{F}(s)] d\tilde{\mathbb{P}} &= \int_A Y d\tilde{\mathbb{P}} = \tilde{\mathbb{E}}[\mathbb{1}_A Y] \\ \int_A \frac{1}{Z(s)} \mathbb{E}[YZ(t)|\mathcal{F}(s)] d\tilde{\mathbb{P}} &= \tilde{\mathbb{E}}[\mathbb{1}_A \frac{1}{Z(s)} \mathbb{E}[YZ(t)|\mathcal{F}(s)]] \\ &= \tilde{\mathbb{E}}[\mathbb{E}[\mathbb{1}_A \frac{1}{Z(s)} YZ(t)|\mathcal{F}(s)]] \text{ use Lemma 5.2.1} \\ &= \mathbb{E}[Z(s) \mathbb{E}[\mathbb{1}_A \frac{1}{Z(s)} YZ(t)|\mathcal{F}(s)]] \\ &= \mathbb{E}[\mathbb{E}[\mathbb{1}_A YZ(t)|\mathcal{F}(s)]] = \mathbb{E}[\mathbb{1}_A YZ(t)] \\ \text{(using again Lemma 5.2.1)} &= \tilde{\mathbb{E}}[\mathbb{1}_A Y] = \int_A Y d\tilde{\mathbb{P}} \\ \text{So, } \frac{1}{Z(s)} \mathbb{E}[YZ(t)|\mathcal{F}(s)] &= \tilde{\mathbb{E}}[Y|\mathcal{F}(s)] \end{aligned}$$

Let $\{W(t) : 0 \leq t \leq T\}$ be a Brownian motion on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, and let $\{\mathcal{F}(t) : 0 \leq t \leq T\}$ be its natural filtration, and assume $\mathcal{F} = \mathcal{F}(T)$.

Consider a stock with price process $\{S(t) : 0 \leq t \leq T\}$ with

$$S(t) = S(0) \exp \left\{ \int_0^t e^{-u} dW(u) + \int_0^t \left(1 - \frac{1}{2} e^{-2u}\right) du \right\}$$

a. Let

$$X(t) = \int_0^t e^{-u} dW(u) + \int_0^t \left(1 - \frac{1}{2} e^{-2u}\right) du$$

Determine the distribution of $X(t)$.

b. Prove the $\{S(t) : t \geq 0\}$ is an Itô process.

Answer 6a. Let $Y(t) = \int_0^t e^{-u} dW(u)$. Since $Y(t)$ is the Itô integral of a deterministic process, by Theorem 4.4.9 $Y(t)$ is normally distributed with $\mathbb{E}[Y(t)] = 0$ and $\text{Var}[Y(t)] = \int_0^t e^{-2u} du = \frac{1}{2}(1 - e^{-2t})$. Since $X(t) = Y(t) + \int_0^t (1 - \frac{1}{2} e^{-2u}) du = Y(t) + t + \frac{1}{4}(e^{-2t} - 1)$, we see that $X(t)$ is normally distributed with mean $\mathbb{E}[X(t)] = t + \frac{1}{4}(e^{-2t} - 1)$ and variance $\text{Var}[X(t)] = \text{Var}[Y(t)] = \frac{1}{2}(1 - e^{-2t})$.

Answer 6b. With $X(t) = \int_0^t e^{-u} dW(u) + \int_0^t (1 - \frac{1}{2}e^{-2u}) du$ we have $dX(t) = e^{-t} dW(t) + (1 - \frac{1}{2}e^{-2t}) dt$ and $dX(t)dX(t) = e^{-2t} dt$. Note that $S(t) = S(0)e^{X(t)}$, so let $f(x) = S(0)e^x$, then $f_x(x) = f_{xx}(x) = f(x)$. By the Itô-Doebelin formula, we have,

$$\begin{aligned} dS(t) &= (X(t)) = S(t)dX(t) + \frac{1}{2}S(t)dX(t)dX(t) \\ &= S(t) \left(e^{-t} dW(t) + (1 - \frac{1}{2}e^{-2t}) dt \right) + \frac{1}{2}S(t)e^{-2t} dt \\ &= S(t)dt + S(t)e^{-t} dW(t). \end{aligned}$$

This shows that $S(t) = S(0) + \int_0^t S(u)du + \int_0^t S(u)e^{-u}dW(u)$. Hence, $\{S(t) : t \geq 0\}$ is an Itô process.